Moderate and extreme urban rainfall modeling at a fine spatio-temporal resolution

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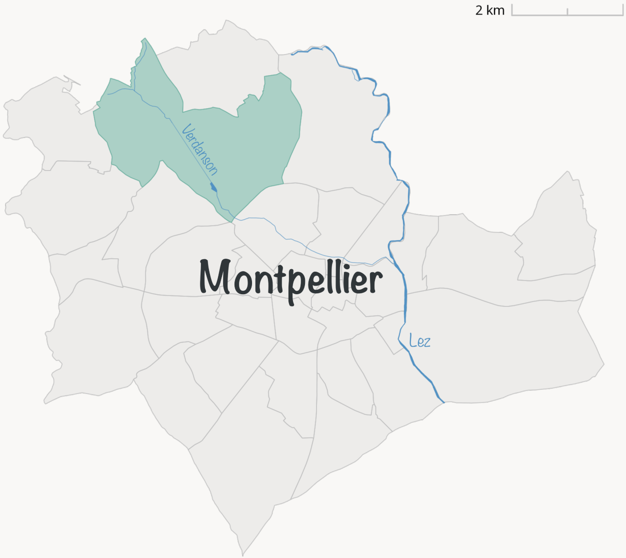
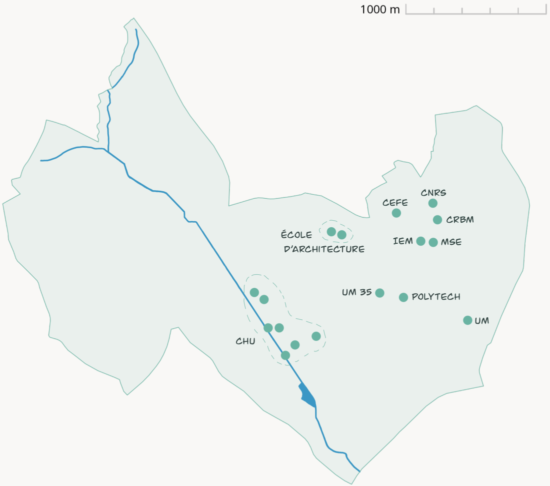
Study area



- **Geography:**
  Verdanson water catchment, tributary of the Lez, located in an urban area
- **Context:**
  Mediterranean events, flood risks

Data

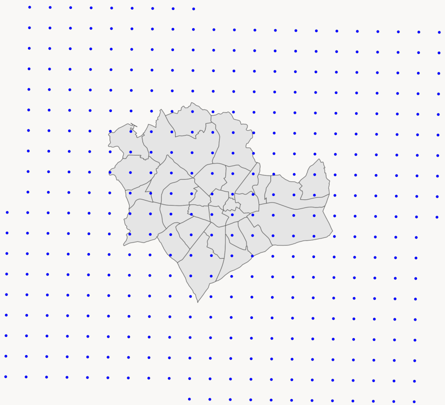


- **Source:** Urban observatory of HydroScience Montpellier (OHSM)[1]
- **Time period:** $[2019, 2022]$
- **High temporal resolution:**
  Every minute $\longrightarrow$ 5-minute aggregation
- **High spatial resolution:**
  Interdistance $\in [77, 1531]$ meters

[1] Finaud-Guyot et al. 2023

$\mathcal{S}$ = {**17 rain gauges**} $\subset \mathbb{R}^2$ and $\mathcal{T} \subset \mathbb{R}_+$

Additional data: COMEPHORE



- **Source:** Météo France
- **Time period:** [1997, 2023]
- **Temporal resolution:** Every hour
- **Spatial resolution:** 1 km $\times$ 1 km

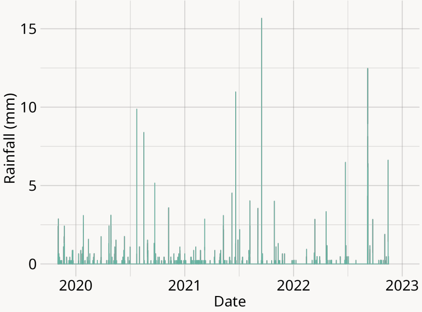
$\mathcal{S} = \{\textbf{400 pixels}\} \subset \mathbb{R}^2$ and $\mathcal{T} \subset \mathbb{R}_+$

Generalized Pareto Distribution

$$\overline{H}_\xi\left(\frac{x-u}{\sigma}\right) = \begin{cases} \left(1+\xi\frac{x-u}{\sigma}\right)_+^{-1/\xi} & \text{if } \xi \neq 0\,, \\ e^{-\frac{x-u}{\sigma}} & \text{if } \xi = 0\,, \end{cases}$$

where $a_+ = \max(a, 0)$, $\sigma > 0$, $x - u > 0$

- Models extreme precipitation
- Depends on a threshold choice



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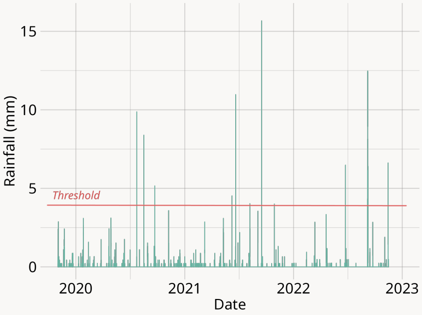

Generalized Pareto Distribution → **Extended GPD**[1]

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where $a_+ = \max(a,0)$, $\sigma > 0$, $x-u > 0$

- Models extreme precipitation
- Depends on a threshold choice

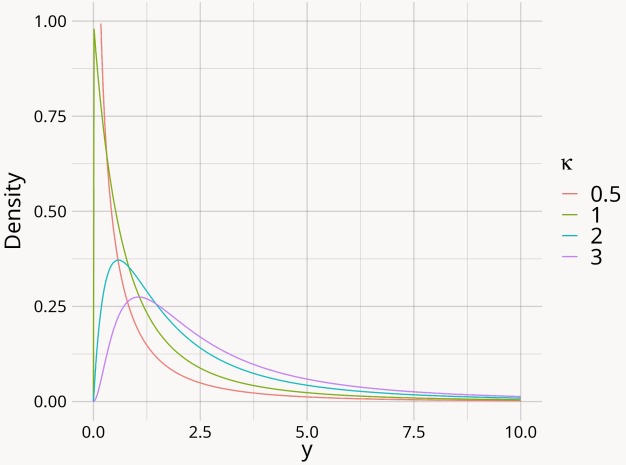
$$F(x) = G\left(H_\xi\left(\frac{x}{\sigma}\right)\right),$$

where $G(x) = x^\kappa$, $\kappa > 0$

- Models moderate and extreme precipitation
- Avoid a threshold choice

[1] Naveau et al. 2016

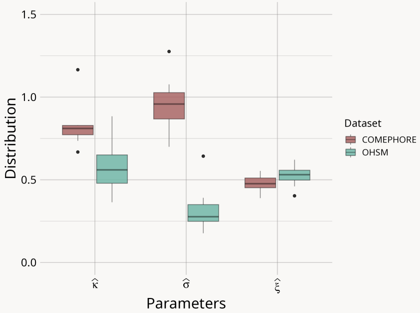
Generalized Pareto Distribution → **Extended GPD**



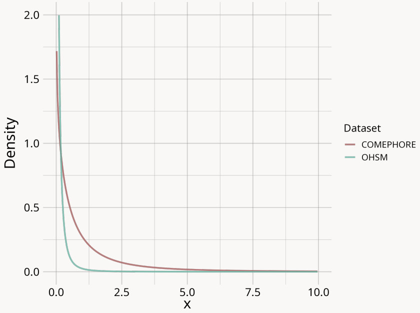
$\sigma = 1$, $\xi = 0.5$

EGPD fitting

Left-censoring: selected according to the NRMSE criterion for each site individually



Parameter estimates



Density with mean parameters

Rainfall field: $\boldsymbol{X} = \{X_{\boldsymbol{s},t}\,,\ (\boldsymbol{s},t) \in \mathcal{S} \times \mathcal{T}\}$

Assumptions: $\boldsymbol{X}$ is a stationary isotropic max-stable Brown-Resnick process

Brown-Resnick process (Brown and Resnick 1977)

For all $\boldsymbol{s} \in \mathcal{S}$ and $t \in \mathcal{T}$,

$$X_{\boldsymbol{s},t} = \bigvee_{j=1}^{\infty} \xi_j e^{W^j_{\boldsymbol{s},t} - \gamma(\boldsymbol{s},t)}$$

- ξ_j: point of a Poisson process with intensity $\xi^{-2} d\xi$
- W^j: independent replicates of an intrinsic stationary and isotropic Gaussian random field $\boldsymbol{W}$
- γ: spatio-temporal variogram of $\boldsymbol{W}$

Let $\Lambda_{\mathcal{S}} \subset \mathbb{R}_+^2$ and $\Lambda_{\mathcal{T}} \subset \mathbb{R}_+$ be sets of spatial and temporal lags respectively.

Spatio-temporal extremogram

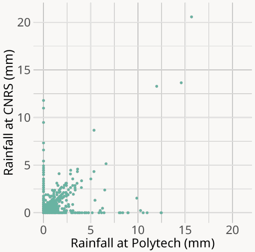
For all $\boldsymbol{h} \in \Lambda_{\mathcal{S}}, \tau \in \Lambda_{\mathcal{T}}$,

$$\chi(\boldsymbol{h}, \tau) = \lim_{q \to 1} \chi_q(\boldsymbol{h}, \tau), \qquad \text{where} \qquad \chi_q(\boldsymbol{h}, \tau) = \mathbb{P}(X^*_{\boldsymbol{s},t} > q \mid X^*_{\boldsymbol{s}+\boldsymbol{h},t+\tau} > q),$$

with $q \in [0, 1[$ and $X^*_{\boldsymbol{s},t}$ the standardized univariate margins.

Spatio-temporal variogram γ

$$\gamma(\boldsymbol{h}, \tau) = \frac{1}{2} Var\left(W_{\boldsymbol{s},t} - W_{\boldsymbol{s}+\boldsymbol{h},t+\tau}\right), \; \boldsymbol{h} \in \Lambda_{\mathcal{S}}, \tau \in \Lambda_{\mathcal{T}}$$



Dependence measures

Let $\Lambda_{\mathcal{S}} \subset \mathbb{R}_+^2$ and $\Lambda_{\mathcal{T}} \subset \mathbb{R}_+$ be sets of spatial and temporal lags respectively.

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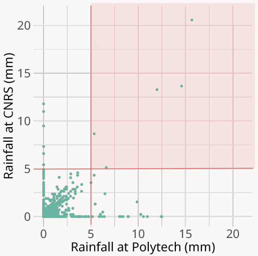

Spatio-temporal extremogram of a Brown-Resnick process

Let $\boldsymbol{h} \in \Lambda_{\mathcal{S}}$ and $\tau \in \Lambda_{\mathcal{T}}$. We have

$$\chi(\boldsymbol{h}, \tau) = 2\left(1 - \phi\left(\sqrt{\frac{1}{2}\gamma(\boldsymbol{h}, \tau)}\right)\right)$$

with ϕ the std normal c.d.f. and γ the variogram of $\boldsymbol{W}$.

Dependence model framework: Buhl et al. 2019

Case of additive separability: $\frac{\gamma(\boldsymbol{h},\tau)}{2} = \beta_1 ||\boldsymbol{h}||^{\alpha_1} + \beta_2 \tau^{\alpha_2}, \; 0 < \alpha_1, \alpha_2 \leq 2\,, \; \beta_1, \beta_2 > 0$

Spatio-temporal

$$\chi(\boldsymbol{h},\tau) = 2\left(1 - \phi\left(\sqrt{\tfrac{1}{2}\gamma(\boldsymbol{h},\tau)}\right)\right)$$

Transformation:

$$\eta(\chi) = 2\log\left(\phi^{-1}\left(1 - \tfrac{1}{2}\chi\right)\right)$$

Spatial

$$\eta\left(\chi(\boldsymbol{h},0)\right) = \log\beta_1 + \alpha_1 \log||\boldsymbol{h}||\,, \;\; \boldsymbol{h} > 0$$

Temporal

$$\eta\left(\chi(0,\tau)\right) = \log\beta_2 + \alpha_2 \log\tau\,, \;\; \tau > 0$$

Spatio-temporal dependence modeling

Case of additive separability: $\frac{\gamma(\boldsymbol{h},\tau)}{2} = \beta_1 ||\boldsymbol{h}||^{\alpha_1} + \beta_2 \tau^{\alpha_2}, \; 0 < \alpha_1, \alpha_2 \leq 2\,, \; \beta_1, \beta_2 > 0$

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Spatial

$$\eta\left(\chi(\boldsymbol{h},0)\right) := c_1 + \alpha_1 x_h\,, \; h > 0$$

Temporal

$$\eta\left(\chi(0,\tau)\right) := c_2 + \alpha_2 x_\tau\,, \; \tau > 0$$

Weighted Least Squares Estimation (WLSE)

$$\begin{pmatrix} \widehat{c}_i \\ \widehat{\alpha}_i \end{pmatrix} = \text{argmin}_{c_i,\alpha_i} \sum_x w_x \left(\eta\left(\widehat{\chi}\right) - (c_i + \alpha_i x)\right)^2$$

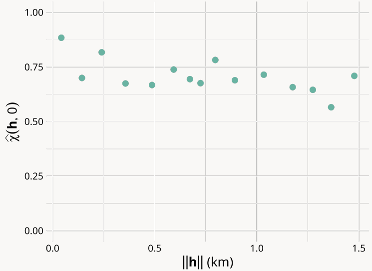
Spatial dependence estimation

Empirical spatial extremogram

For a fixed $t \in \mathcal{T}$ and q a high quantile,

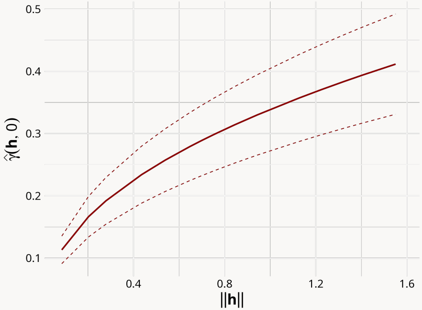
$$\widehat{\chi}_q^{(t)}(\boldsymbol{h}, 0) = \frac{\frac{1}{|N_{\boldsymbol{h}}|} \sum_{i,j \,|\, (\boldsymbol{s}_i, \boldsymbol{s}_j) \in N_{\boldsymbol{h}}} \mathbb{1}_{\{X^*_{\boldsymbol{s}_i,t} > q \,, X^*_{\boldsymbol{s}_j,t} > q\}}}{\frac{1}{|\mathcal{S}|} \sum_{i=1}^{|\mathcal{S}|} \mathbb{1}_{\{X^*_{\boldsymbol{s}_i,t} > q\}}},$$

where $C_{\boldsymbol{h}}$ are equifrequent distance classes and $N_{\boldsymbol{h}} = \left\{(\boldsymbol{s}_i, \boldsymbol{s}_j) \in \mathcal{S}^2 \;\middle|\; \|\boldsymbol{s}_i - \boldsymbol{s}_j\| \in C_{\boldsymbol{h}}\right\}$.



Transformation and WLSE

$\Longrightarrow$

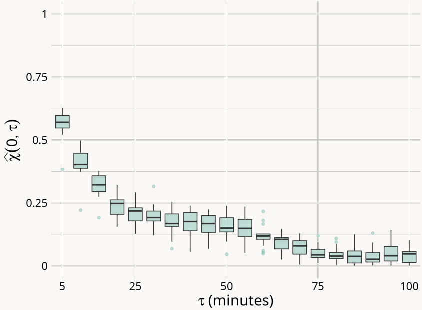


Spatial variogram $\widehat{\gamma}(\boldsymbol{h}, 0) = 2\widehat{\beta}_1 \|\boldsymbol{h}\|^{\widehat{\alpha}_1}$

Empirical temporal extremogram

For a location $s \in \mathcal{S}$, a high quantile q and $t_k \in \{t_1, \dots, t_T\}$,

$$\widehat{\chi}_q^{(s)}(0,\tau) = \frac{\frac{1}{T-\tau}\sum_{k=1}^{T-\tau} \mathbb{1}_{\{X^*_{s,t_k}>q\,,X^*_{s,t_k+\tau}>q\}}}{\frac{1}{T}\sum_{k=1}^{T} \mathbb{1}_{\{X^*_{s,t_k}>q\}}}$$



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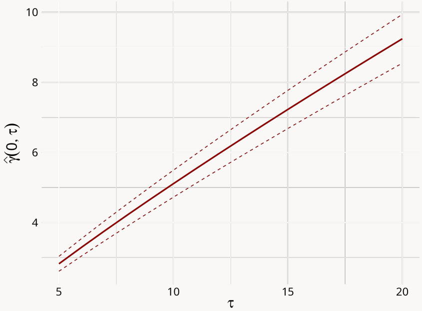


Temporal variogram $\widehat{\gamma}(0,\tau) = 2\widehat{\beta}_2\tau^{\widehat{\alpha}_2}$

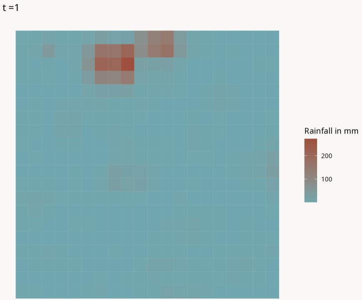
Brown-Resnick simulations

- Spatial: $\mathcal{S} = \{\text{ 400 sites }\}$, $\mathcal{T} = \{1, \ldots, 50\}$, $|\Lambda_{\mathcal{S}}| = 10$ and $|\Lambda_{\mathcal{T}}| = 10$
- Temporal: $\mathcal{S} = \{\text{ 25 sites }\}$, $\mathcal{T} = \{1, \ldots, 300\}$, $|\Lambda_{\mathcal{S}}| = 10$ and $|\Lambda_{\mathcal{T}}| = 10$

	True	Mean	RMSE	MAE
$\widehat{\beta}_1$	0.4	0.524	0.138	0.126
$\widehat{\alpha}_1$	1.5	1.507	0.120	0.088
$\widehat{\beta}_2$	0.2	0.259	0.093	0.074
$\widehat{\alpha}_2$	1	0.873	0.149	0.128

Parameter estimates for 100 realisations

True variogram: $\frac{1}{2}\gamma(\boldsymbol{h}, \tau) = 0.4\|\boldsymbol{h}\|^{3/2} + 0.2\tau$



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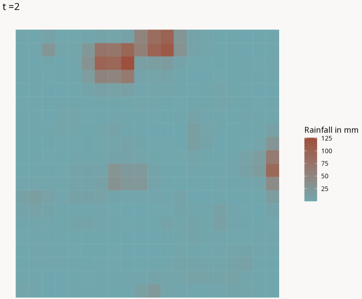

Model validation

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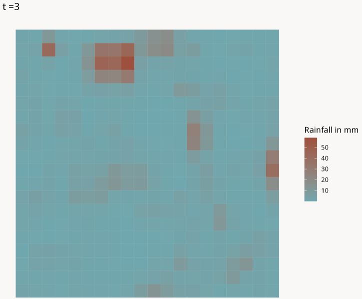

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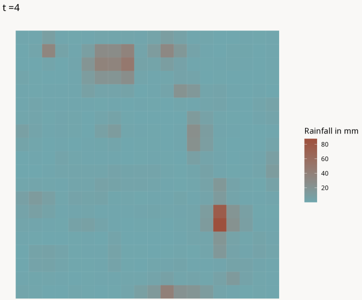

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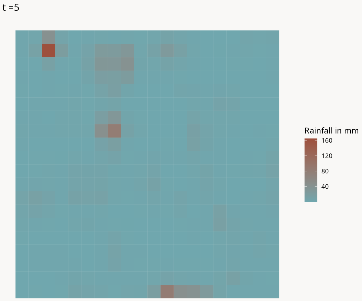

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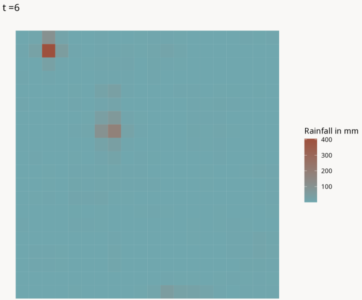

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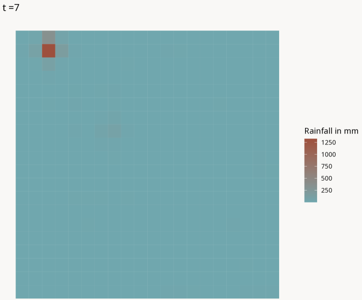

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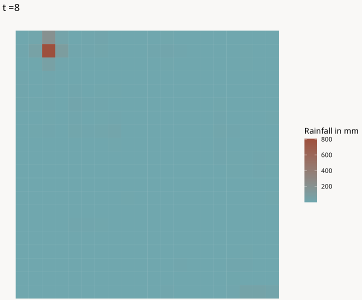

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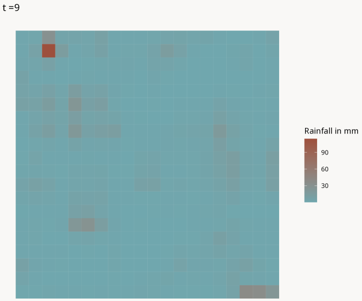

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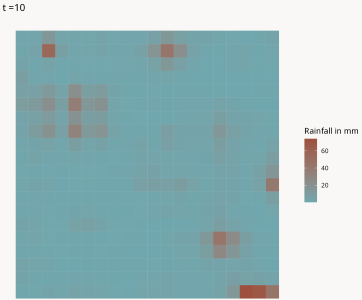
True variogram: $\frac{1}{2}\gamma\left(\boldsymbol{h}, \tau\right) = 0.4\|\boldsymbol{h}\|^{3/2} + 0.2\tau$



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- Spatial: $\mathcal{S} = \{\text{ 400 sites }\}$, $\mathcal{T} = \{1, ..., 50\}$, $|\Lambda_{\mathcal{S}}| = 10$ and $|\Lambda_{\mathcal{T}}| = 10$
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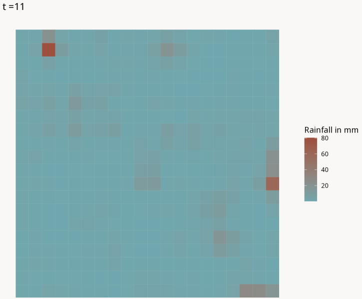

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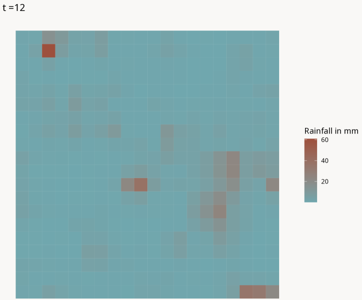

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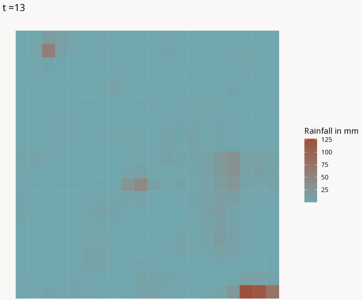

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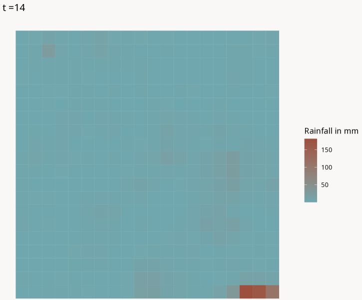

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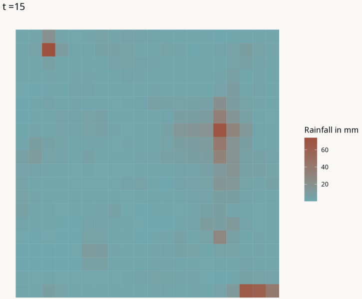

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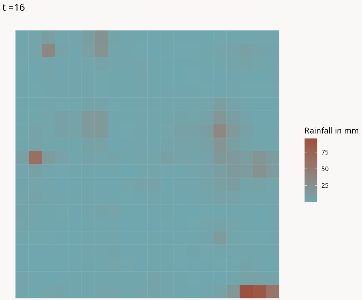

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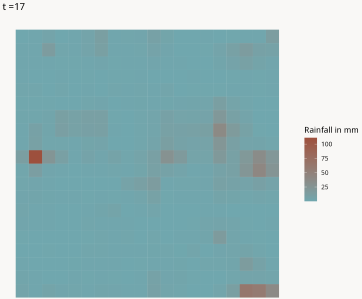

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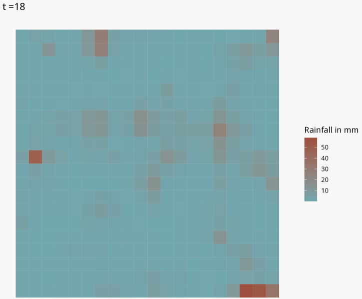

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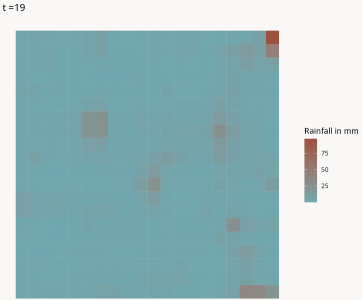

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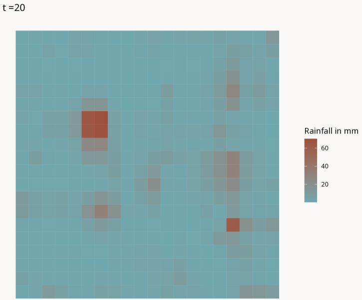

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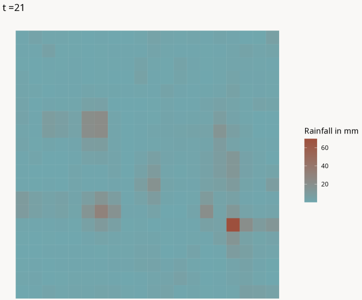

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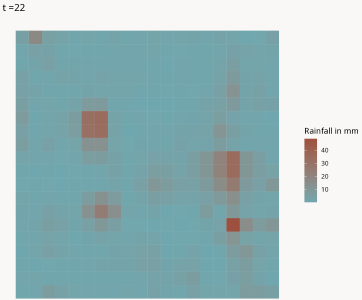

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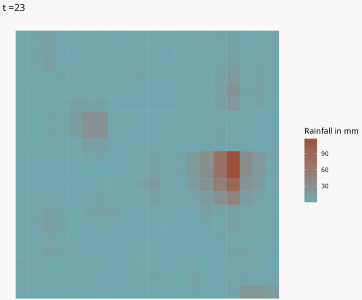

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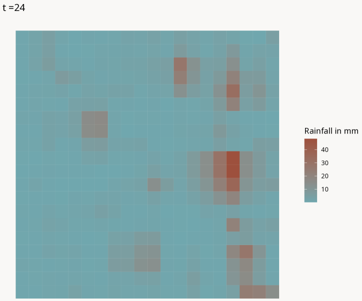

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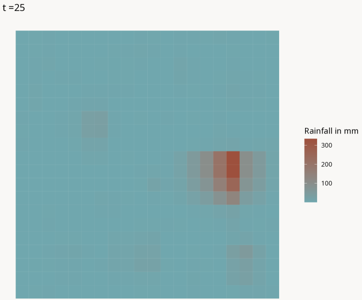

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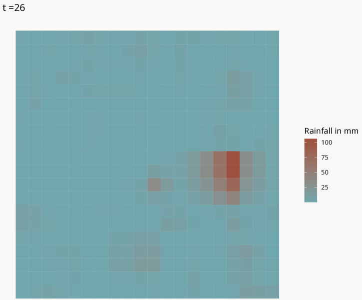

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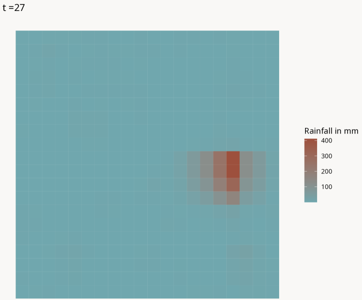

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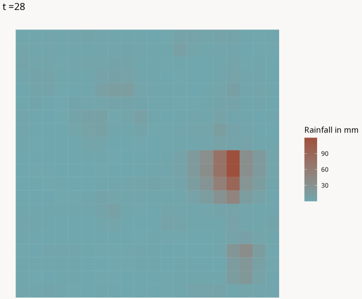

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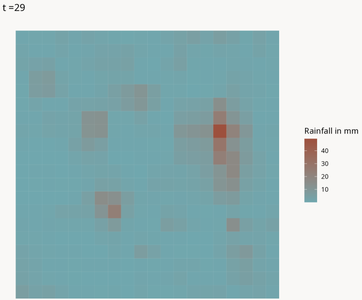

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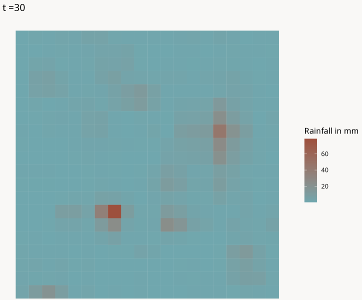

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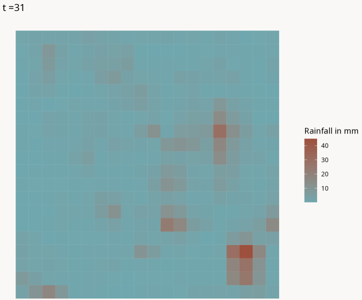

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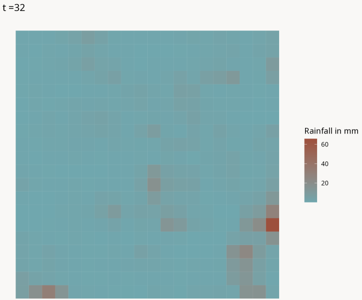

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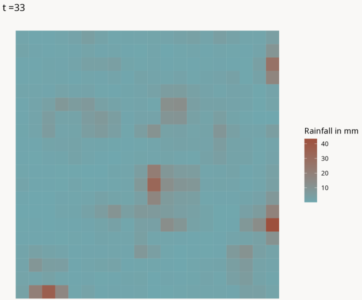

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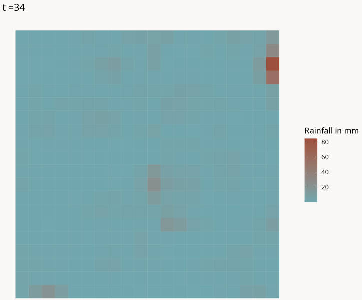

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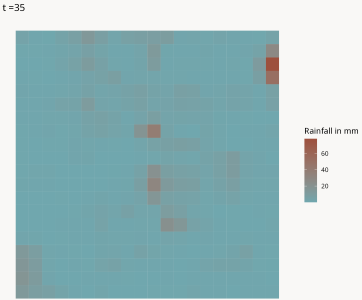

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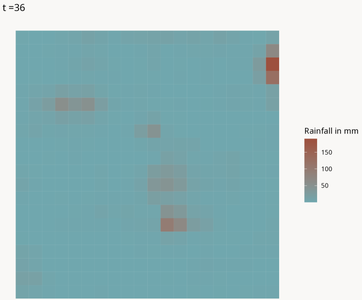

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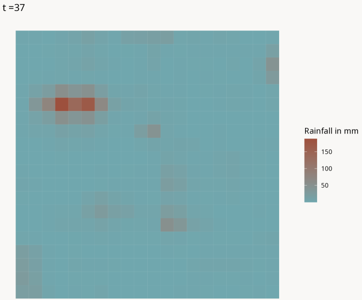

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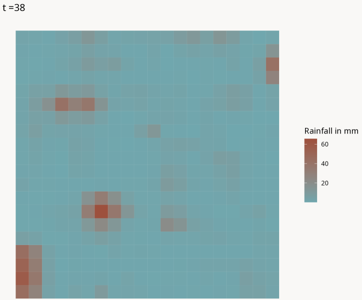

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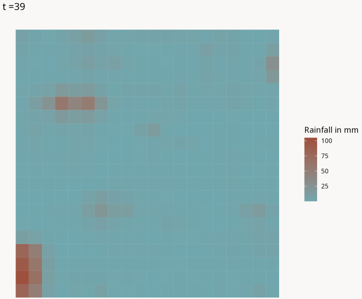

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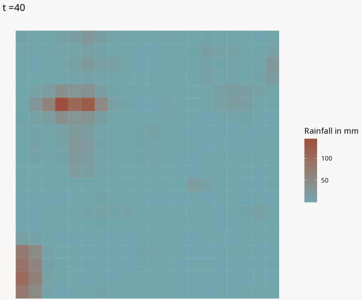

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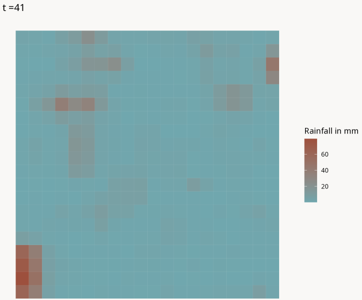

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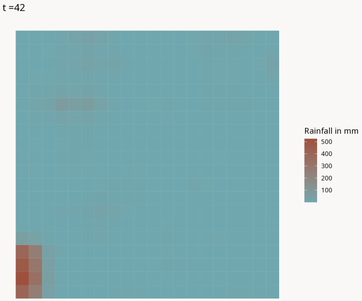

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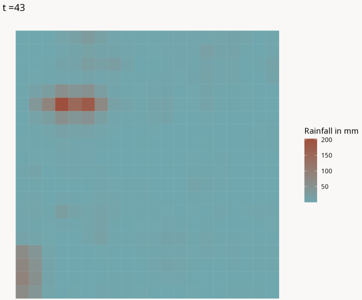

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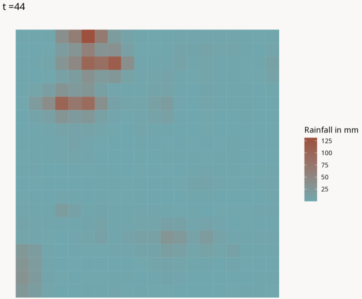

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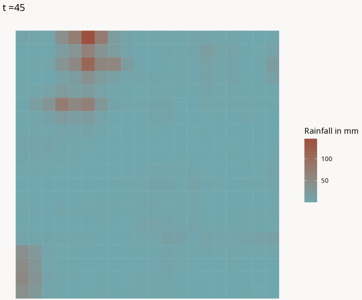

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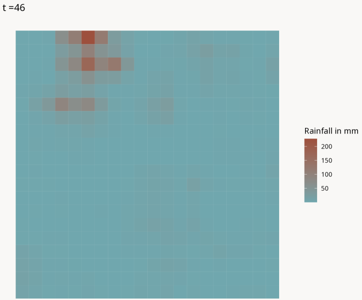

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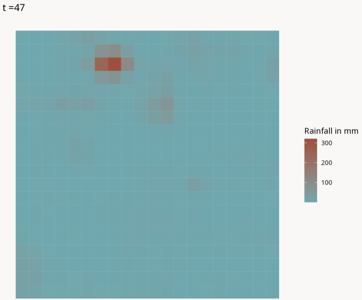

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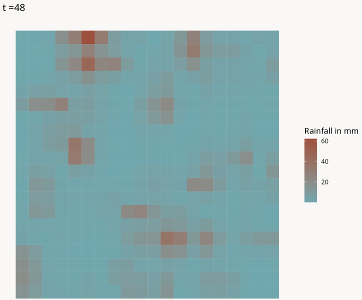

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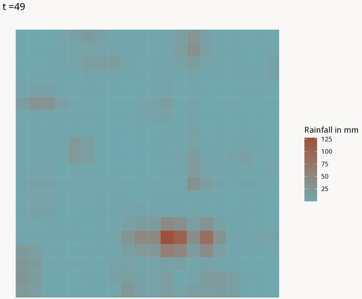

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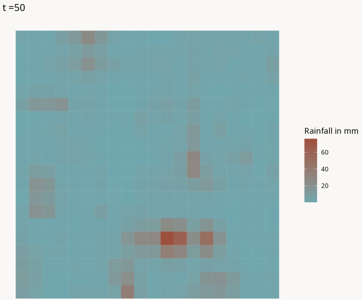
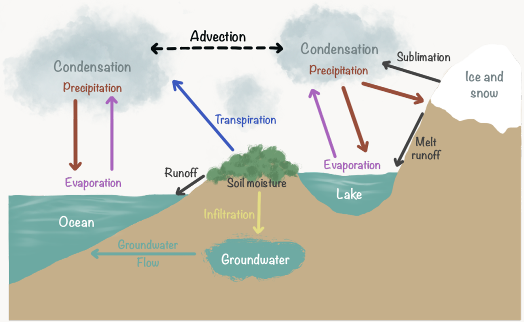


Advection vector $\boldsymbol{V}$

- Horizontal transport of air masses
- To relax the separability assumption



Hydrologic cycle

Considering advection

Advection vector $\boldsymbol{V}$

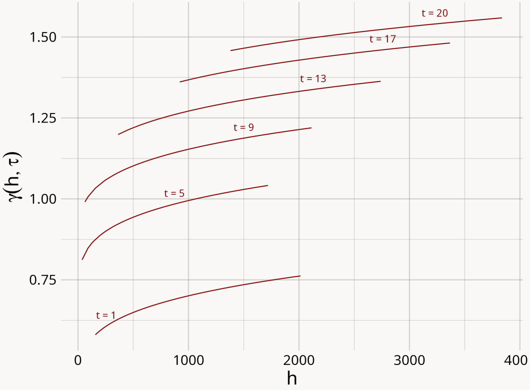
- Horizontal transport of air masses
- To relax the separability assumption

Lagrangian/Eulerian variogram

$$\gamma_L(\boldsymbol{h}, \tau) = \gamma(\boldsymbol{h} - \tau\boldsymbol{V}, \tau)$$

Dependence model

$$\frac{1}{2}\gamma_L(\boldsymbol{h}, \tau) = \beta_1 \|\boldsymbol{h} - \tau\boldsymbol{V}\|^{\alpha_1} + \beta_2 \tau^{\alpha_2}$$



Spatial variogram with a constant advection $\boldsymbol{V} = (0.001, 45)^T$ on OHSM data

Parameter optimization of $\boldsymbol{\Theta} = (\beta_1, \beta_2, \alpha_1, \alpha_2, \boldsymbol{V})$

Excesses: for p a spatio-temporal configuration

$$E_p = \mathbb{1}_{\{X^*_{\mathbf{s}_i,t_i} > q \,|\, X^*_{\mathbf{s}_j,t_j} > q\}} \sim \mathcal{B}(\chi_{p,\boldsymbol{\Theta}}) \implies \sum_{k=1}^{n_p} E_{p,\,k} \sim \mathcal{B}(n_p\,;\,\chi_{p,\boldsymbol{\Theta}})$$

Composite log-likelihood:

$$\log\left(L_{\boldsymbol{\Theta}}\left(\underline{E}\right)\right) = \sum_p \left[\log\binom{n_p}{k_p} + k_p \log \chi_{p,\boldsymbol{\Theta}} + (n_p - k_p)\log(1 - \chi_{p,\boldsymbol{\Theta}})\right]$$

Perspectives

- Advection estimation on COMEPHORE data
- Incorporing advection in the dependence model
- Downscaling
- Adding wind data

References

Brown, Bruce M. and Sidney I. Resnick (1977). "Extreme values of independent stochastic processes". In: *Journal of Applied Probability*. DOI: 10.2307/3213346.

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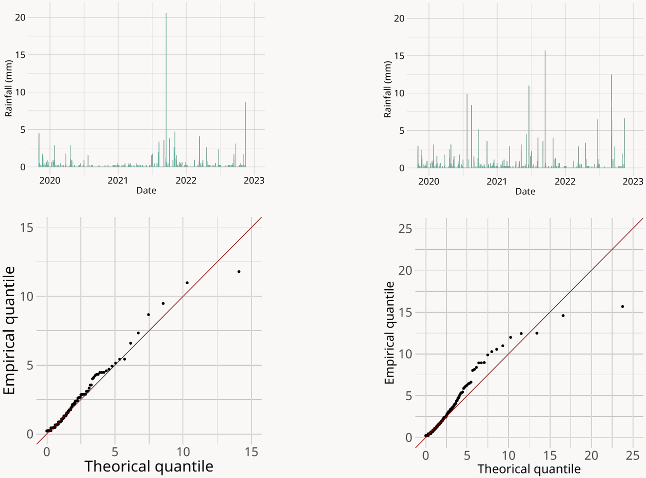
Finaud-Guyot, Pascal et al. (2023). *Rainfall data collected by the HSM urban observatory (OMSEV)*. DOI: 10.23708/67LC36.

Naveau, Philippe et al. (2016). "Modeling jointly low, moderate, and heavy rainfall intensities without a threshold selection". In: *Water Resources Research*. DOI: 10.1002/2015WR018552.

EGPD fitting



EGPD fitting on CNRS and Polytech rain gauges
($\widehat{\kappa} = 0.56, \widehat{\sigma} = 0.26$ et $\widehat{\xi} = 0.51$)